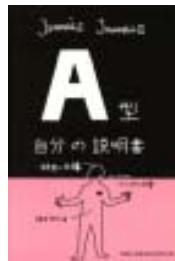


Likelihood-based method for estimating penetrance and disease allele frequency

Yuki Sugaya and Ritei Shibata
Keio University, Japan

Topic in Japan



BUNGEISYA CO., LTD.

- 5 millions seller
- Blood type character analysis
- ¥1,050 yen (in tax)

Blood type character analysis in Japan

- Type A
 - polite, cooperative, responsible, oversensitive, and somewhat uptight
- Type B
 - optimistic, social, and easygoing to the point of occasional laziness
- Type O
 - active, strong-willed, friendly, and blunt
- Type AB
 - moody, eccentric, withdrawn, and prone to bouts of brilliance and/or insanity

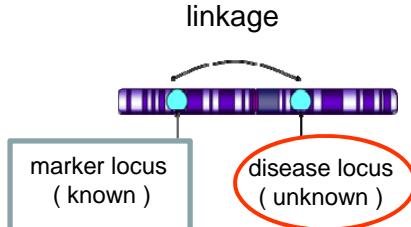
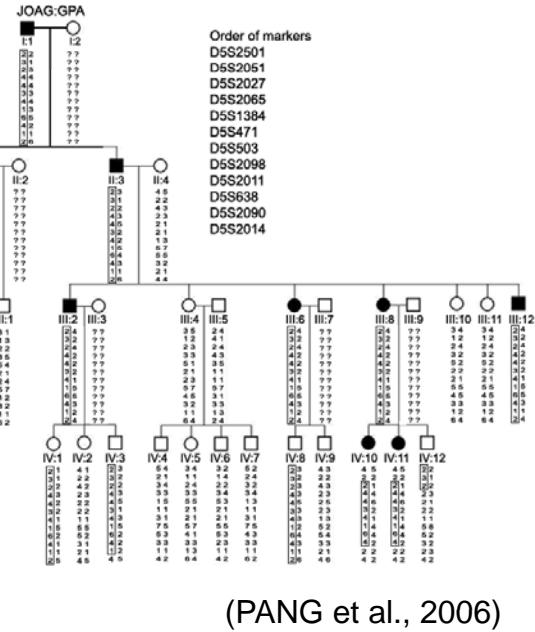
Blood type

phenotype	Type A	Type B	Type O	Type AB
genotype	$A/A \ A/O$	$B/B \ B/O$	O/O	A/B

$$\left. \begin{array}{l} P(\text{TypeA} \mid A/A) = 1 \\ P(\text{TypeB} \mid A/A) = 0 \\ \vdots \end{array} \right\} \text{penetrance}$$

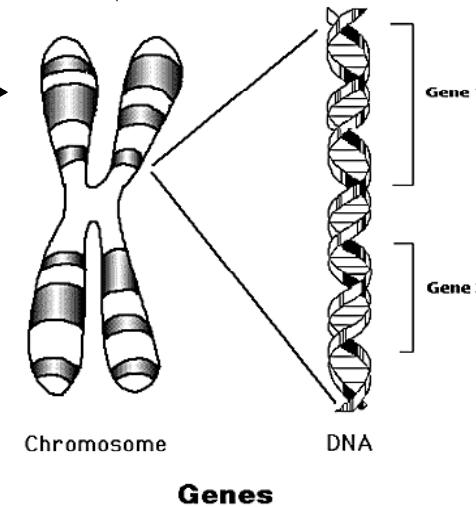
Linkage analysis

pedigree data



where?

disease
locus →

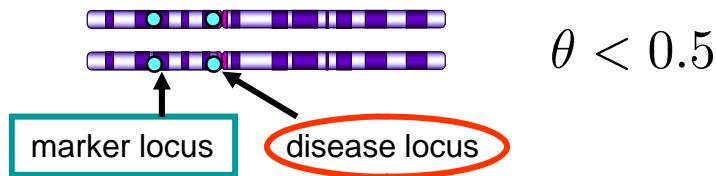


Objective

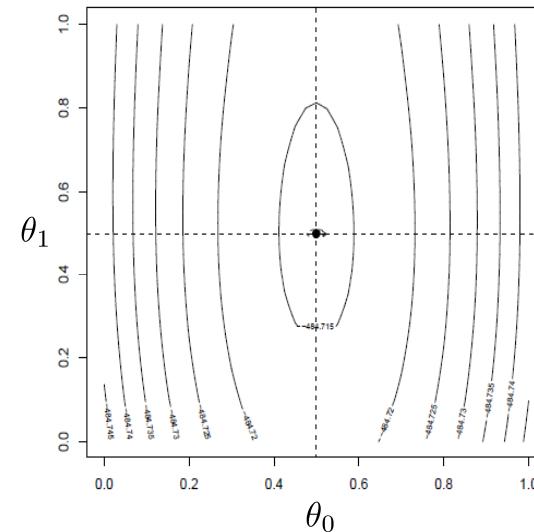
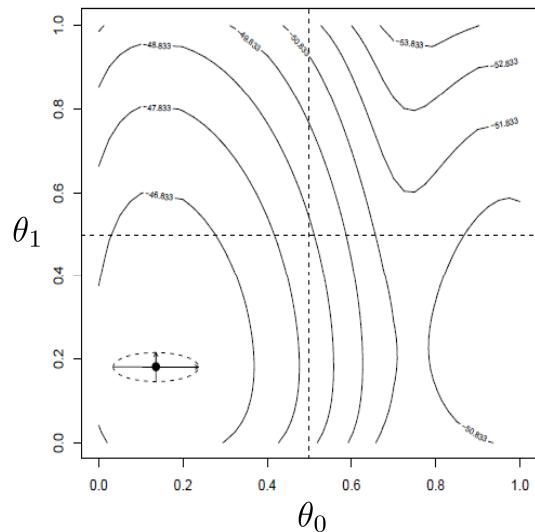
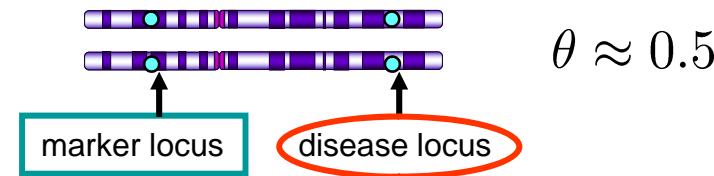
To detect a gene associated with a disease
from pedigree data

Recombination fraction

- Two loci are close to



- Two loci are far from

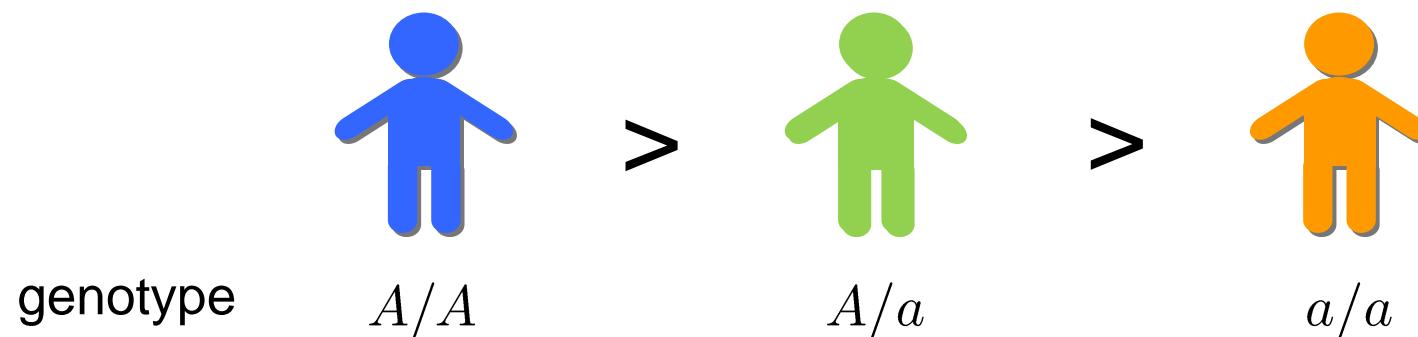


We detect the disease gene through recombination fraction.

Assumption

- The disease is caused by a pair of disease susceptibility allele A

a : normal allele



Penetrance

- Conditional probability of observing the corresponding phenotype (affected status) given the specified genotype



$$P(\text{Affected} \mid A/A) = \alpha$$

A/A



$$P(\text{Affected} \mid A/a) = \beta$$

A/a



$$P(\text{Affected} \mid a/a) = \gamma$$

a/a

$$\alpha > \beta > \gamma$$

Normal linkage analysis

- Penetrance
 - $\alpha = 1, \beta = 1, \gamma = 0 \Rightarrow$ dominant model
 - $\alpha = 1, \beta = 0, \gamma = 0 \Rightarrow$ recessive model

⋮

- Disease susceptibility allele frequency f_A

$$f_A = 0.0001, 0.001, 0.1 \dots$$

There are a lot of choices of parameters.
Different assumption may lead us to different result.

Previous works

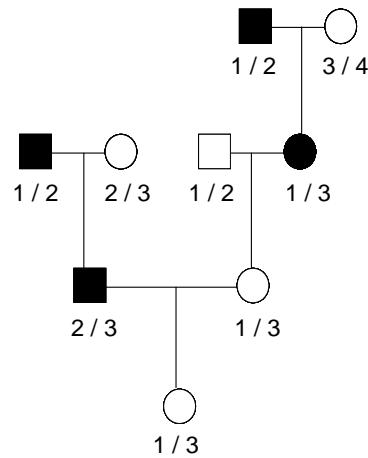
- Swart et al. (2004)
 - penetrance and allele frequency
 - $\alpha = \beta$ or $\beta = \gamma$

- Wang et al. (2006)
 - penetrance
 - nuclear family

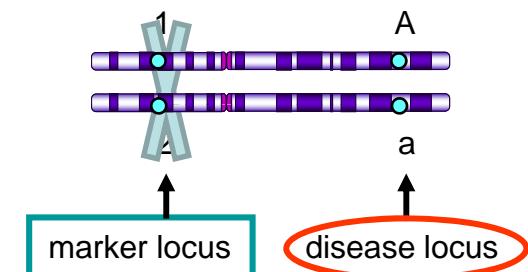
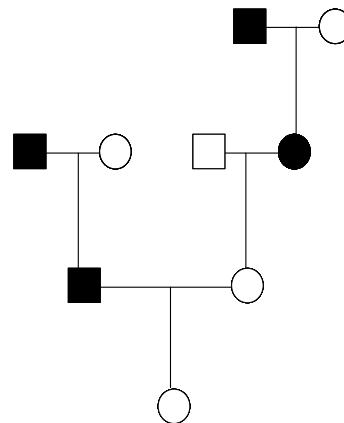
Likelihood of pedigree data

linkage analysis

$$P(\mathcal{A}(V) \cap \mathcal{M}(V))$$



$$P(\mathcal{A}(V))$$



A_v : affected status

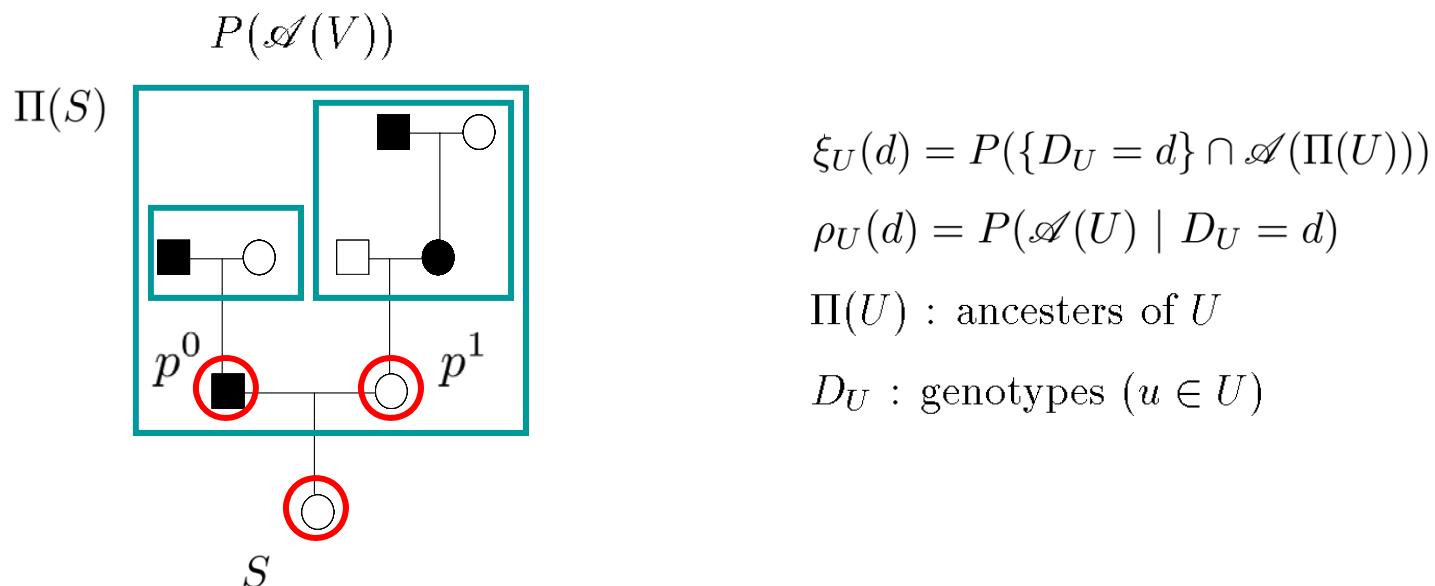
M_v : marker genotypes

marker information is not necessary

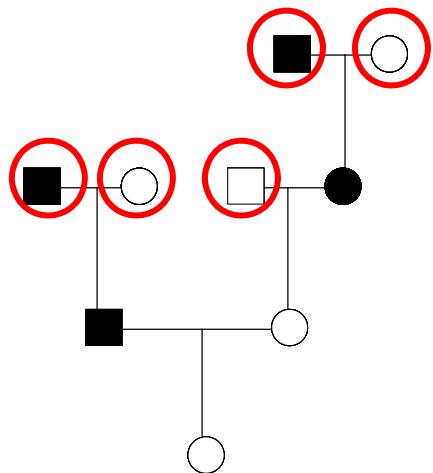
Calculation of the likelihood (1)

$$\xi_S(d) = \sum_{(d_0, d_1)} \frac{\rho_p((d_0, d_1)) P(D_S = d \mid D_p = (d_0, d_1)) \xi_{p^0}(d_0) \xi_{p^1}(d_1)}{\text{penetrance} \quad \text{inheritance probability}}$$

$$\begin{cases} \alpha, 1 - \alpha, \\ \beta, 1 - \beta, \\ \gamma, 1 - \gamma \end{cases} \quad (0, \frac{1}{2}, 1)$$



Calculation of the likelihood (2)



v : founder

$$\xi_v(d) = P(\{D_v = d\} \cap \underline{\mathcal{A}(\Pi(v))})$$

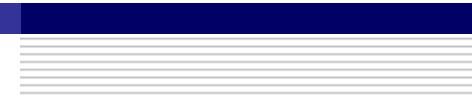
empty set



$$\xi_v(d) = P(\{D_v = d\})$$
$$= \underline{P(\{D_v^0 = d_0\})} \underline{P(\{D_v^1 = d_1\})}$$

disease susceptibility allele frequency
($f_A, 1 - f_A$)

Parameter estimation

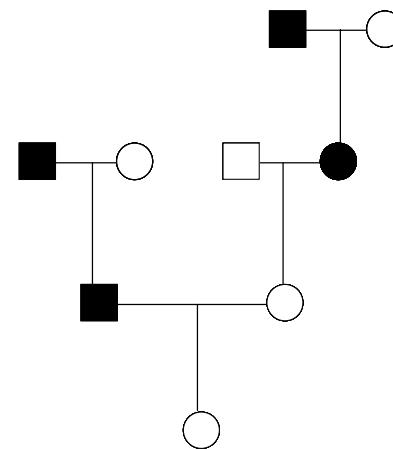


$$P(\mathcal{A}(V)) = L(\alpha, \beta, \gamma, f_A) \xrightarrow{\quad} \max_{\alpha, \beta, \gamma, f_A}$$

$$L(\alpha, \beta, \gamma, f_A) = \sum c_{ijkl} \alpha^i \beta^j \gamma^k f_A^l$$

$\max(i + j + k)$: informative individuals

$\max(l)$: 2 × founders



Optimization problem

$$\text{maximize} \quad L(\mathbf{x}) \quad \mathbf{x} = (\alpha, \beta, \gamma, f_A)$$

subject to $\gamma \leq \beta \leq \alpha, 0 \leq \alpha \leq 1, 0 \leq \beta \leq 1,$
 $0 \leq \gamma \leq 1$ and $0 \leq f_A \leq 1$

$$\leftrightarrow \left[\begin{array}{l} \alpha - \beta \geq 0 : g_1(\mathbf{x}) \\ \beta - \gamma \geq 0 : g_2(\mathbf{x}) \\ \alpha \geq 0 : g_3(\mathbf{x}) \\ \beta \geq 0 : g_4(\mathbf{x}) \\ \gamma \geq 0 : g_5(\mathbf{x}) \\ f_A \geq 0 : g_6(\mathbf{x}) \\ \alpha \geq 0 : g_7(\mathbf{x}) \\ \beta \geq 0 : g_8(\mathbf{x}) \\ \gamma \geq 0 : g_9(\mathbf{x}) \\ 1 - f_A \geq 0 : g_{10}(\mathbf{x}) \end{array} \right]$$

POAG pedigree data (PANG et al., 2006)

Primary Open Angle Glaucoma

30 pedigree members

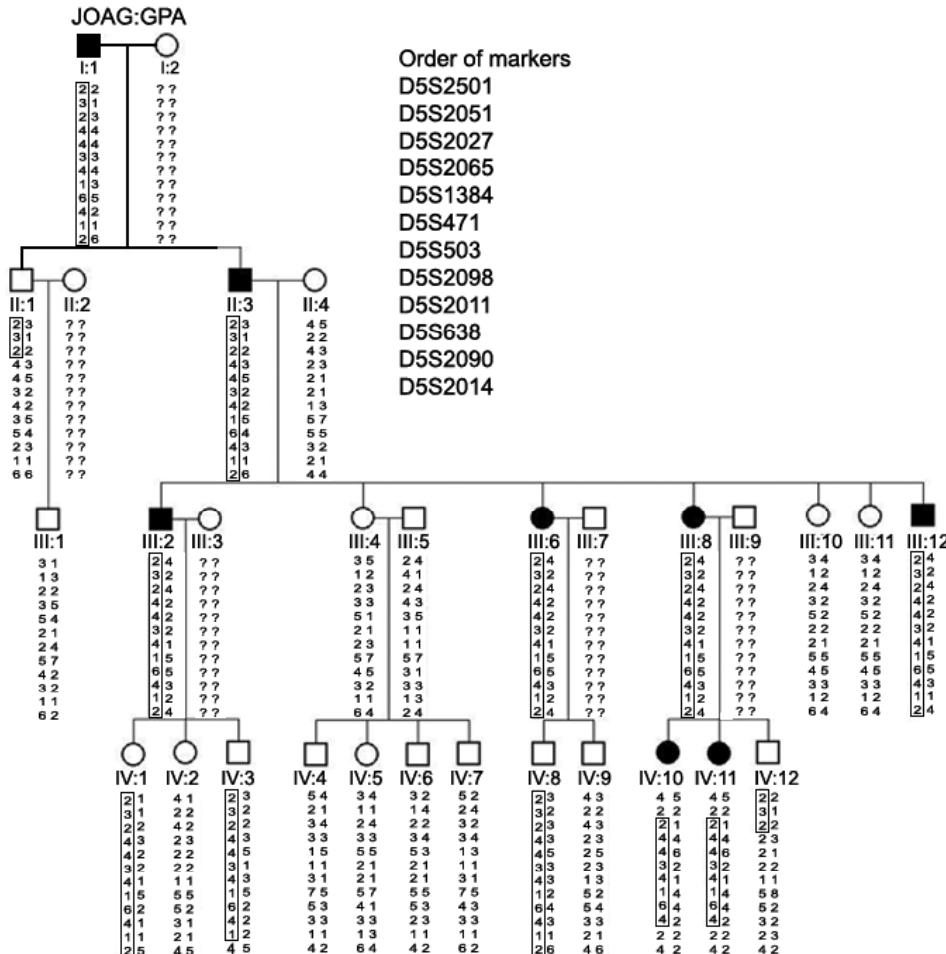
- { 8 affected individuals
- 22 non-affected individuals

PANG et al. assume

$$(\alpha, \beta, \gamma) = (1, 1, 0)$$

(dominant model)

$$f_A = 0.0001$$



Maximum likelihood estimates

Kuhn-Tucker Conditions

$\hat{\mathbf{x}}$: local maximum

$$\Rightarrow \exists \mu_i \leq 0 \quad (i = 1, \dots, 10)$$

$$\nabla L(\hat{\mathbf{x}}) + \sum_{i=1}^{10} \mu_i \nabla g_i(\hat{\mathbf{x}}) = \mathbf{0}$$

$$\mu_i g_i(\hat{\mathbf{x}}) = 0 \quad (i = 1, \dots, 10)$$

$$\hat{\mathbf{x}} = (0.921, 0.921, 0, 0.065)$$

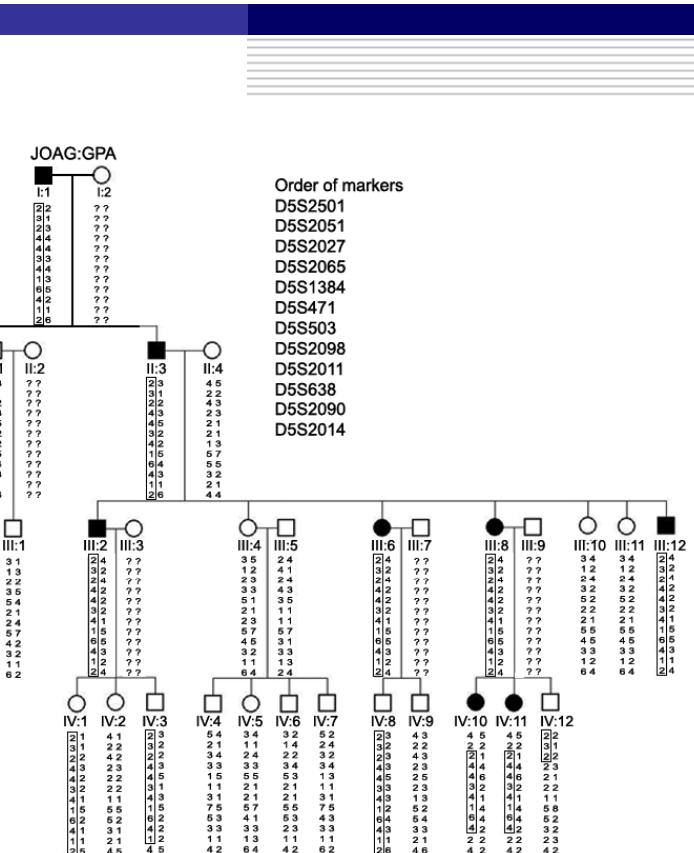
$$\mu_1 = -1.510 \times 10^{-2}$$

$$\mu_5 = -7.351$$

$$\mu_2 = \mu_3 = \mu_4 = \mu_6 = \mu_7 = \mu_8 = \mu_9 = \mu_{10} = 0$$

$\hat{\mathbf{x}}$ satisfy Kuhn-Tucker condition

$$\nabla L(\hat{\mathbf{x}}) = (-1.510 \times 10^{-2}, 1.510 \times 10^{-2}, -7.351, 7.064 \times 10^{-5})$$



Probability inheritance algorithm (Sugaya and Shibata)

- Probability inheritance algorithm
 - polynomial form

$$L(\boldsymbol{\theta}) = \sum_{i,j=0}^k \alpha_{ij} \theta_0^i \theta_1^j$$

$$\boldsymbol{\theta} = (\theta_0, \theta_1)$$

θ_0 : paternal recombination fraction

θ_1 : maternal recombination fraction

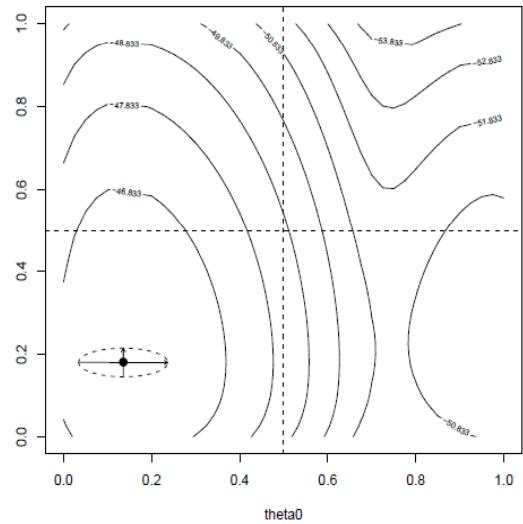
k : non-founder

- Elston-Stewart algorithm (1971)
 - backward
 - large pedigree, small number of markers
 - Linkage package
- Lander-Green algorithm (1987)
 - inheritance vector
 - hidden markov model
 - small pedigree, large number of markers
 - Genehunter

$$\theta_0 \neq \theta_1$$

✗ polynomial form

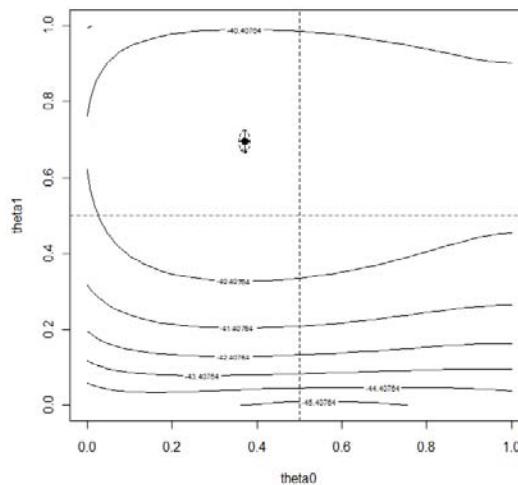
Visual interpretation



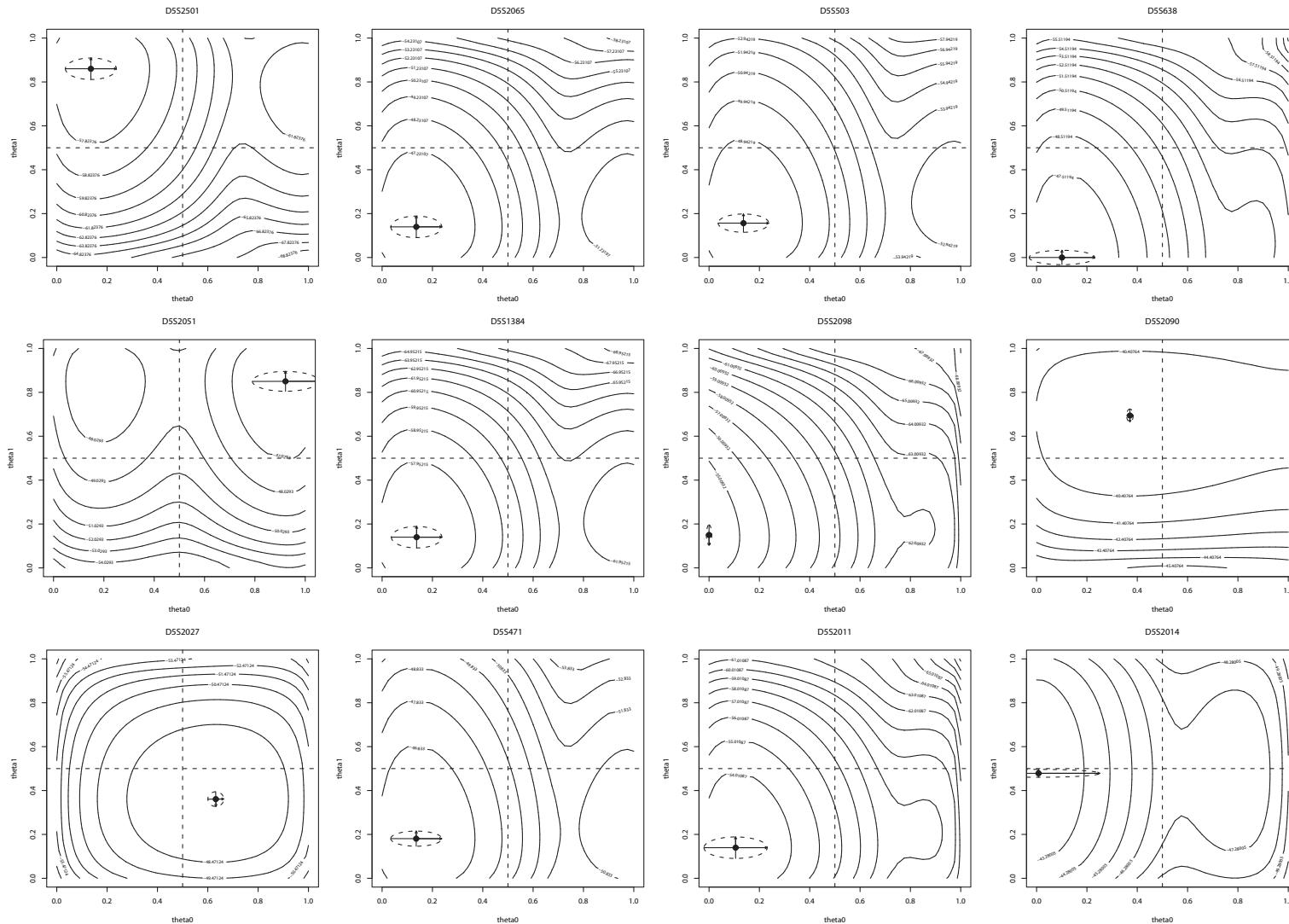
- Contour plot
- Fisher information matrix

$$I(\boldsymbol{\theta}) = E\left(-\frac{\partial^2}{\partial \boldsymbol{\theta} \partial \boldsymbol{\theta}^T} \log L(\boldsymbol{\theta})\right)$$

$$\rightarrow -\frac{\partial^2}{\partial \boldsymbol{\theta} \partial \boldsymbol{\theta}^T} \log L(\boldsymbol{\theta})|_{\boldsymbol{\theta}=\hat{\boldsymbol{\theta}}}$$



- $\theta_i > 0.5$
- Fisher information is low \Rightarrow NA



recombination fraction	D5S2501	D5S2051	D5S2027	D5S2065	D5S1384	D5S471	D5S563	D5S2098	D5S2011	D5S638	D5S2090	D5S2014
θ_0	0.136	NA	NA	0.136	0.137	0.136	0.137	NA	0.106	0.101	NA	0.008
θ_1	NA	NA	NA	0.140	0.140	0.181	0.157	NA	0.140	0	NA	NA

→
disease locus?

Comparison to dominant model

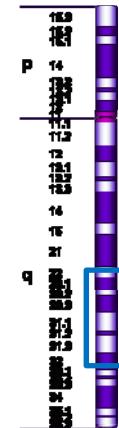
marker	dominant model		estimated values	
	$(\hat{\theta}_0, \hat{\theta}_1)$	log likelihood	$(\hat{\theta}_0, \hat{\theta}_1)$	log likelihood
D5S2501	(0.200, NA)	-62.787	(0.136, NA)	-56.824
D5S2051	(NA, NA)	-51.858	(NA, NA)	-46.029
D5S2027	(NA, NA)	-53.023	(NA, NA)	-47.471
D5S2065	(0.200, 0.200)	-52.195	(0.136, 0.140)	-46.231
D5S1384	(0.200, 0.200)	-62.887	(0.137, 0.140)	-56.952
D5S471	(0.200, 0.239)	-51.766	(0.136, 0.181)	-45.833
D5S503	(0.200, 0.216)	-53.861	(0.137, 0.157)	-47.942
D5S2098	(NA, NA)	-59.411	(NA, NA)	-54.009
D5S2011	(0.174, 0.200)	-59.047	(0.106, 0.140)	-53.011
D5S638	(0.174, 0.067)	-52.556	(0.101, 0)	-46.512
D5S2090	(NA, NA)	-44.993	(NA, NA)	-39.408
D5S2014	(0.085, NA)	-48.257	(0.008, NA)	-42.280

Linkage analysis by PANG et al.

- Linkage Package (MLINK)
- dominant model
- disease allele susceptibility frequency : 0.0001
- 5q22.1 – q32
- $\theta_0 = \theta_1$

$$\log \frac{L(\theta)}{L(1/2)}$$

Chromosome 5



Marker	position(Mb)	0	0.01	0.05	0.1	0.2	0.3	0.4	θ max
D5S2501	110.1	-Inf	-7.52	-3.49	-1.9	-0.55	-0.02	0.13	0.4
D5S2051	111	-Inf	-7.08	-3.55	-2.08	-0.79	-0.26	-0.1	0.4
D5S2027	-	-Inf	-7.05	-3.64	-2.25	-1	-0.4	-0.1	0.4
D5S2065	113.7	-Inf	-1.54	0.35	0.97	1.26	1.09	0.66	0.2
D5S1384	118.9	-Inf	-1.54	0.34	0.97	1.26	1.09	0.66	0.2
D5S471	119.1	-Inf	-1.81	0.09	0.72	1.05	0.93	0.57	0.2
D5S503	120.3	-Inf	-1.67	0.21	0.84	1.15	1	0.61	0.2
D5S2098	-	-Inf	1.6	2.06	2.07	1.78	1.33	0.74	0.1
D5S2011	141.3	-Inf	-1.24	0.6	1.18	1.39	1.15	0.68	0.2
D5S638	146.7	-Inf	-0.12	1.12	1.48	1.53	1.22	0.7	0.2
D5S2090	147.2	-Inf	-5.78	-3.05	-1.93	-0.91	-0.41	-0.13	0.4
D5S2014	149.9	-Inf	-0.56	0.7	1.1	1.2	0.96	0.52	0.2